Errata, Mahler Study Aids for MAS2, 2022 HCM, 6/6/22 Page 1

In June 2022, the CAS announced that effective with the Fall 2022 exam sittings, the guessing penalty for exams MAS-I and MAS-II will be eliminated. **Therefore, you should make sure to choose a letter response for every question.** 

Read the portion of my second study guide on Bayes Analysis prior to the McElreath textbook.

1, page 6:

1 - Z is sometimes referred to as the "complement of credibility".

Confusingly, often instead what I have denoted Y, the item that is given weight 1 - Z, is referred to as the "complement of credibility".

Tse in <u>Nonlife Actuarial Models: Theory Methods and Evaluation</u> does not use the term "complement of credibility". Weight 1 - Z is given to M = manual rate.

2, page 208, last line: = 0.4259.

**2**, solution 6.3:  $(24)(1/\omega)(1/\omega)(1/\omega)(1/\omega) = 24/\omega^4$ . Final solution is okay.

**2**, page 306, 5th line from the bottom: (300)<sup>2</sup>(0.21)

**3**, solution 5.1:  $f(x) = \{(a+b-1)! / (a-1)! (b-1)!\} (x/\theta)^{a-1} \{1 - (x/\theta)\}^{b-1} / \theta =$ 

**3**, Q. 6.45: a Beta Distribution with  $\theta = 1$ 

**3**, solution 7.46: a' = a + # successes = 9 + 15 = 24, and b' = b + # failures = 5 + 25 = 30. Posterior mean is: a' / (a' + b') = 24 / (24 + 30) = 0.444.

**3**, Q. 13.12: What is the posterior density function for the parameter  $\boldsymbol{\theta}$  for this insured?

**3**, Solution: 13.28: Thus at x = 1, 8, 27, 125, the densities are proportional to:  $\beta^{-1/3} \exp(-\beta^{-1/3})$ ,  $\beta^{-1/3} \exp(-2\beta^{-1/3})$ ,  $\beta^{-1/3} \exp(-3\beta^{-1/3})$ , and  $\beta^{-1/3} \exp(-5\beta^{-1/3})$ .

3, page 275, solutions 8.16 and 8.17 are missing:

**8.16.** E. A Gamma-Exponential with prior Gamma with  $\alpha = 1$  and  $\theta = 1/100$ . Thus the posterior Gamma has parameters  $\alpha = 1 + 2 = 3$ , and  $1/\theta = 100 + 40 + 80 = 220$ .  $E[1/\lambda] =$  negative first moment of the posterior Gamma = 220/(3-1) = 110. Alternately, the posterior distribution is proportional to:  $e^{-100\lambda}\lambda e^{-40\lambda} \lambda e^{-80\lambda} = \lambda^2 e^{-220\lambda}$ .  $E[X \mid \lambda] = 1/\lambda$ . Therefore, the expected size of the next claim is:

$$\frac{\int\limits_{0}^{\infty} (1/\lambda) \lambda^2 e^{220\lambda} d\lambda}{\int\limits_{0}^{\infty} \lambda^2 e^{220\lambda} d\lambda} = \frac{220^2 \Gamma(2)}{220^3 \Gamma(3)} = 220/2 = 110.$$

<u>Comment</u>: Since  $\alpha = 1 \le 2$ , one can not apply Buhlmann Credibility. The posterior mixed distribution is Pareto with  $\alpha = 3$  and  $\theta = 220$ , with mean 220/(3-1) = 110.

**8.17. D.** A Gamma-Exponential.

Thus the posterior Gamma has parameters  $\alpha = 4 + 3 = 7$ , and  $1/\theta = 1000 + 100 + 200 + 500 = 1800$ .  $E[1/\lambda] =$  negative first moment of the posterior Gamma = 1800/(7-1) = 300. Alternately,  $K = \alpha - 1 = 4 - 1 = 3$ . Z = (3/(3+K) = 3/(3+3) = 50%. The prior mean is the negative first moment of the prior Gamma = 1000/(4-1) = 1000/3. Estimate of the future severity is: (50%)(800/3) + (50%)(1000/3) = 300. Comment: The posterior mixed distribution is Pareto with  $\alpha = 7$  and  $\theta = 1800$ , with mean 1800/(7-1) = 300. **4**, solution 4.39: Z = **150**/(**150**+0.986) = 0.9935.

Estimated future pure premium is: (0.9935)(3130.20) + (1 - 0.9935)(2783.49) = 3128. Estimated aggregate claim amount for the next period: (3128)(175) = 547,400.

5, Q. 2.10: Model III: K-Nearest Neighbors with K = 5.

**5**, page 88, third paragraph from the bottom: Thus the **Gini Index** for a region is: (2)(proportion of yeses)(proportion of nos).

**5**, Q.4.3: For observation 6, LoyalCH = 0.695. (For given LoyalCH = 0.795, the prediction would be CH.)

5, solution. 8.16: Final solution is okay.

 $Z_1$  and  $Z_3$  are uncorrelated. It turns out that constraining  $Z_3$  to be uncorrelated with  $Z_1$  is equivalent to constraining the direction  $\phi_3$  to be orthogonal (perpendicular)

to the direction  $\phi_1$ . Thus,  $\sum_{j=1}^5 \phi_{j1} \phi_{j3} = 0$ . Thus Statement III is true.

<u>Comment</u>:  $Z_{\underline{1}}$ ,  $Z_{\underline{2}}$ , and  $Z_{\underline{3}}$  are each orthogonal to the others.

**5**, Q. 10.13:  $x_1 = (-1, 0) \quad x_3 = (2, -1),$ 

**5**, Solution 10.1: {56} is 21 from {7**7**} Then on the second page of the solution. {23, 39} is {(56 - 23) + (56 - 39) + (77 - 23) + (77 - 39)}/4 = 35.5 away from {56, 7**7**}. Final solution is okay.